

# Testing Block Sphericity of a Covariance Matrix

*Prueba de Esfericidad por Bloques  
para una Matriz de Covarianza*

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## Abstract

This article deals with the problem of testing the hypothesis that  $q$   $p$ -variate normal distributions are independent and that their covariance matrices are equal. The exact null distribution of the likelihood ratio statistic when  $q = 2$  is obtained using inverse Mellin transform and the definition of Meijer's  $G$ -function. Results for  $p = 2, 3, 4$  and  $5$  are given in computable series forms.

**Key words and phrases:** block sphericity, distribution, inverse Mellin transform, Meijer's  $G$ -function, residue theorem.

## Resumen

Este artículo trata el problema de probar la hipótesis de que  $q$  distribuciones normales  $p$ -variadas son independientes y con matrices de covarianza son iguales. La distribución exacta nula del estadístico de razón de verosimilitudes cuando  $q = 2$  se obtiene usando la transformada inversa de Mellin y la definición de la  $G$ -función de Meijer. Los resultados para  $p = 2, 3, 4$  y  $5$  se dan en forma de series calculables.

**Palabras y frases clave:** esfericidad por bloques, distribución, transformada inversa de Mellin, función  $G$  de Meijer, teorema del residuo.

## 1 Introduction

Suppose that the  $pq \times 1$  random vector  $\mathbf{X}$  has a multivariate normal distribution with mean vector  $\boldsymbol{\mu}$  and covariance matrix  $\Sigma$  and that  $\mathbf{X}$ ,  $\boldsymbol{\mu}$  and  $\Sigma$  are partitioned as

$$\mathbf{X} = (\mathbf{X}'_1 \quad \mathbf{X}'_2 \quad \cdots \quad \mathbf{X}'_q)', \quad \boldsymbol{\mu} = (\boldsymbol{\mu}'_1 \quad \boldsymbol{\mu}'_2 \quad \cdots \quad \boldsymbol{\mu}'_q)'$$

and

$$\Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} & \cdots & \Sigma_{1q} \\ \Sigma_{21} & \Sigma_{22} & \cdots & \Sigma_{2q} \\ \vdots & & & \\ \Sigma_{q1} & \Sigma_{q2} & \cdots & \Sigma_{qq} \end{pmatrix}$$

where  $\mathbf{X}_i$  and  $\boldsymbol{\mu}_i$  are  $p \times 1$  and  $\Sigma_{ij}$  is  $p \times p$ ,  $i, j = 1, \dots, q$ . Consider testing the null hypothesis  $H$  that the subvectors  $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_q$  are independent and that the covariance matrices of these subvectors are equal, i.e.,

$$H : \Sigma = \begin{pmatrix} \Delta & 0 & \cdots & 0 \\ 0 & \Delta & \cdots & 0 \\ \vdots & & & \\ 0 & 0 & \cdots & \Delta \end{pmatrix}$$

against the alternative  $H_a$  that  $H$  is not true. In  $H$  the common covariance matrix  $\Delta$  is unspecified. Olkin, while extending the circular symmetric model to the case where the symmetries are exhibited in blocks, defined  $H$  and called it *block sphericity hypothesis* (see [13]). Let  $A$  be the sample sum of squares and product matrix formed from a sample of size  $N = n + 1$ . Partition  $A$  as  $A = (A_{ij})$  where  $A_{ij}$  is  $p \times p$ ,  $i, j = 1, \dots, q$ . The likelihood ratio statistic for testing  $H$  (see [17]) is given by

$$\Lambda = \frac{\det(A)^{\frac{1}{2}N}}{\det\left(\frac{1}{q} \sum_{i=1}^q A_{ii}\right)^{\frac{1}{2}qN}}.$$

The null hypothesis is rejected if  $\Lambda \leq \Lambda_0$  where  $\Lambda_0$  is determined by the null distribution and level of significance. When  $p = 1$ , the hypothesis of block sphericity reduces to the Mauchly sphericity hypothesis  $H : \Sigma = \sigma^2 I_q$  which has been studied extensively, e.g., see Sugira [15, 16], Khatri and Srivastava

[6, 7], Muirhead [10], Gupta [3], Nagar, Jain and Gupta [11] and Amey and Gupta [1]. The hypothesis of block sphericity is useful in multivariate repeated measures designs (see [17]). Since,  $A \sim W_{pq}(n, \Sigma)$ , the  $h^{\text{th}}$  null moment of  $\Lambda$  can be obtained by integrating over Wishart density. We have

$$\begin{aligned} E(\Lambda^h) &= q^{\frac{1}{2}pqNh} c_{pq,n} \int_{A>0} \frac{\det(A)^{\frac{1}{2}Nh}}{\det\left(\sum_{i=1}^q A_{ii}\right)^{\frac{1}{2}qNh}} \det(A)^{\frac{1}{2}(n-pq-1)} \text{etr}\left(-\frac{1}{2}\Sigma^{-1}A\right) dA \\ &= q^{\frac{1}{2}pqNh} c_{pq,n} \int_{A>0} \det\left(\sum_{i=1}^q A_{ii}\right)^{-\frac{1}{2}qNh} \det(A)^{\frac{1}{2}(Nh+n-pq-1)} \text{etr}\left(-\frac{1}{2}\Sigma^{-1}A\right) dA \end{aligned}$$

where

$$c_{m,n} = \left[ 2^{\frac{1}{2}nm} \pi^{\frac{1}{4}m(m-1)} \prod_{j=1}^m \left[ \frac{1}{2}(n-j+1) \right] \det(\Sigma)^{\frac{1}{2}n} \right]^{-1}. \quad (1.1)$$

Consequently,

$$E(\Lambda^h) = q^{\frac{1}{2}pqNh} \frac{c_{pq,n}}{c_{pq,Nh+n}} E \left[ \det\left(\sum_{i=1}^q A_{ii}\right)^{-\frac{1}{2}qNh} \right] \quad (1.2)$$

where  $A \sim W_{pq}(n+Nh, \Sigma)$ . Under  $H$ ,  $A_{11}, \dots, A_{qq}$  are independent Wishart matrices,  $A_{ii} \sim W_p(n+Nh, \Delta)$ . Hence,  $\sum_{i=1}^q A_{ii} \sim W_p((n+Nh)q, \Delta)$ . Using Theorem 3.3.22 of [5], we obtain

$$\begin{aligned} &E \left[ \det\left(\sum_{i=1}^q A_{ii}\right)^{-\frac{1}{2}qNh} \right] \\ &= 2^{-\frac{1}{2}pqNh} \det(\Delta)^{-\frac{1}{2}qNh} \prod_{i=1}^p \frac{\Gamma[\frac{1}{2}(nq-i+1)]}{\Gamma[\frac{1}{2}(nq+Nqh-i+1)]}. \end{aligned} \quad (1.3)$$

Substituting (1.3) in (1.2) and using (1.1) we obtain the  $h^{\text{th}}$  null moment of  $\Lambda$  as

$$E(\Lambda^h) = q^{\frac{1}{2}pqNh} \prod_{j=1}^{pq} \frac{\Gamma[\frac{1}{2}(n+Nh-j+1)]}{\Gamma[\frac{1}{2}(n-j+1)]} \prod_{i=1}^p \frac{\Gamma[\frac{1}{2}(nq-i+1)]}{\Gamma[\frac{1}{2}(nq+Nqh-i+1)]}. \quad (1.4)$$

The exact non-null distribution of  $\Lambda$ , for  $q = 2$  and under two specific alternatives, has been derived by Gupta and Chao [4]. The asymptotic null

distribution of  $-2 \ln \Lambda$  is chi-square with  $\frac{1}{2}p(q-1)(pq+p+1)$  d.f. The null distribution of  $\Lambda^{\frac{2}{N}}$ , in series involving Bernoulli polynomials, is available in [?, CG]

In this article we will derive the exact null distribution of  $V = \Lambda^{\frac{1}{N}}$  by using inverse Mellin transform and residue theorem (see [14, 12]).

## 2 Exact density of $V$

In order to obtain the exact density function  $f(v)$  of  $V = \Lambda^{\frac{1}{N}}$  we start with the null moment expression. Substituting  $q = 2$  in (1.4) the  $h^{\text{th}}$  moment of  $V$  is simplified as

$$E(V^h) = 2^{ph} \prod_{j=1}^{2p} \left\{ \frac{\Gamma[\frac{1}{2}h + \frac{1}{2}(n-j+1)]}{\Gamma[\frac{1}{2}(n-j+1)]} \right\} \prod_{j=1}^p \left\{ \frac{\Gamma[n - \frac{1}{2}(j-1)]}{\Gamma[h+n - \frac{1}{2}(j-1)]} \right\}$$

This will be rewritten in a slightly different form by using duplication formula for gamma function

$$\Gamma(2z) = \frac{2^{2z-1}}{\sqrt{\pi}} \Gamma(z) \Gamma\left(z + \frac{1}{2}\right).$$

Then

$$E(V^h) = K(n, p) \prod_{j=1}^p \left\{ \frac{\Gamma(h+n-2p-1+2j)}{\Gamma[h+n - \frac{1}{2}(j-1)]} \right\} \quad (2.1)$$

where

$$K(n, p) = \prod_{j=1}^p \left\{ \frac{\Gamma[n - \frac{1}{2}(j-1)]}{\Gamma(n-2p-1+2j)} \right\}. \quad (2.2)$$

Now the density function  $f(v)$  of  $V = \Lambda^{\frac{1}{N}}$  is obtained by taking inverse Mellin transform of  $E(V^h)$  as

$$f(v) = (2\pi\iota)^{-1} \int_L E(V^h) v^{-h-1} dh, \quad 0 < v < 1, \quad (2.3)$$

where  $\iota = \sqrt{-1}$  and  $L$  is a suitable contour. Substituting (2.1) in (2.3) and applying the transformation  $h+n-2p=t$ , one obtains

$$f(v) = K(n, p) v^{n-2p-1} (2\pi\iota)^{-1} \int_{L_1} \prod_{j=1}^p \left\{ \frac{\Gamma(t+2j-1)}{\Gamma[t+2p - \frac{1}{2}(j-1)]} \right\} v^{-t} dt, \quad 0 < v < 1, \quad (2.4)$$

where  $L_1$  is the changed contour and  $K(n, p)$  is defined in (2.2). Now, using the definition of Meijer's  $G$ -function (see [8]), the density of  $V$  is written as

$$f(v) = K(n, p)v^{n-2p-1}G_{p,p}^{p,0} \left[ v \begin{matrix} 2p - \frac{1}{2}(j-1), j=1, \dots, p \\ 2j-1, j=1, \dots, p \end{matrix} \right], 0 < v < 1,$$

where  $G_{p,p}^{p,0}[\ ]$  is the Meijer's  $G$ -function.

Explicit expressions for the density of  $V$  for particular values of  $p$  will be obtained by evaluating the integral in (2.4) with the help of residue theorem.

From (2.4), the density for  $p = 2$  is obtained as

$$f(v) = K(n, 2)v^{n-5}(2\pi i)^{-1} \int_{L_1} \frac{\Gamma(t+1)}{(t+3)\Gamma(t+\frac{7}{2})} v^{-t} dt, 0 < v < 1. \quad (2.5)$$

The integrand has simple poles at  $t = -r - 1$ ,  $r = 0, 1, 3, 4, \dots$ , and a pole of order two at  $t = -3$ . The residue at  $t = -r - 1$ ,  $r \neq 2$ , is

$$\begin{aligned} & \lim_{t \rightarrow -r-1} \left[ \frac{(t+1+r)\Gamma(t+1)}{(t+3)\Gamma(t+\frac{7}{2})} v^{-t} \right] \\ &= \lim_{t \rightarrow -r-1} \left[ \frac{(t+1+r)(t+1+r-1)\cdots(t+1)\Gamma(t+1)}{(t+1+r-1)\cdots(t+1)(t+3)\Gamma(t+\frac{7}{2})} v^{-t} \right] \\ &= \lim_{t \rightarrow -r-1} \left[ \frac{\Gamma(t+1+r+1)}{(t+1+r-1)\cdots(t+1)(t+3)\Gamma(t+\frac{7}{2})} v^{-t} \right] \\ &= \frac{1}{(-1)^{r+1} r! (r-2)\Gamma(\frac{5}{2}-r)} v^{r+1}. \end{aligned}$$

Simplifying  $\Gamma(\frac{5}{2}-r)$  in the above expression using the result  $\Gamma(\beta-j) = \frac{(-1)^j \Gamma(\beta)\Gamma(1-\beta)}{\Gamma(1-\beta+j)}$  we obtain the residue at  $t = -r - 1$  as

$$-\frac{1}{\pi} \frac{\Gamma(r-\frac{3}{2})}{r!(r-2)} v^{r+1}, r \neq 2.$$

The residue at  $t = -3$  is derived as

$$\begin{aligned}
 & \lim_{t \rightarrow -3} \frac{\partial}{\partial t} \left[ \frac{(t+3)^2 \Gamma(t+1)}{(t+3)\Gamma(t+\frac{7}{2})} v^{-t} \right] = \lim_{t \rightarrow -3} \frac{\partial}{\partial t} \left[ \frac{\Gamma(t+4)}{(t+1)(t+2)\Gamma(t+\frac{7}{2})} v^{-t} \right] \\
 &= \lim_{t \rightarrow -3} \left\{ \frac{\partial}{\partial t} \ln \left[ \frac{\Gamma(t+4)}{(t+1)(t+2)\Gamma(t+\frac{7}{2})} \right] - \ln v \right\} \frac{\Gamma(t+4)}{(t+1)(t+2)\Gamma(t+\frac{7}{2})} v^{-t} \\
 &= \lim_{t \rightarrow -3} \left\{ \psi(t+4) - \sum_{i=1,2} (t+i)^{-1} - \psi(t+\frac{7}{2}) - \ln v \right\} \frac{\Gamma(t+4)}{(t+1)(t+2)\Gamma(t+\frac{7}{2})} v^{-t} \\
 &= \left\{ \psi(1) + \frac{1}{2} + 1 - \psi(\frac{1}{2}) - \ln v \right\} \frac{1}{(-2)(-1)\Gamma(\frac{1}{2})} v^3 \\
 &= \left\{ \frac{3}{2} + 2 \ln 2 - \ln v \right\} \frac{1}{2\sqrt{\pi}} v^3
 \end{aligned}$$

where  $\psi(\cdot)$  is the psi function (ver [9]). Now applying the residue theorem to the right hand side of (2.5), the density  $f(v)$  is obtained as

$$f(v) = K(n, 2)v^{n-5} \left[ -\frac{1}{\pi} \sum_{r=0, r \neq 2}^{\infty} \frac{\Gamma(r-\frac{3}{2})}{r!(r-2)} v^{r+1} + \left\{ \frac{3}{2} - \ln\left(\frac{v}{4}\right) \right\} \frac{1}{2\sqrt{\pi}} v^3 \right]$$

for  $0 < v < 1$ . For  $p = 3$ , (2.4) simplifies to

$$f(v) = K(n, 3)v^{n-7}(2\pi\iota)^{-1} \int_{L_1} \frac{\Gamma(t+1)}{(t+5)(t+4)(t+3)\Gamma(t+\frac{11}{2})} v^{-t} dt \quad (2.6)$$

The integrand has simple poles at  $t = -r - 1$ ,  $r = 0, 1, 5, 6, \dots$ , and poles of order two at  $t = -r - 1$ ,  $r = 2, 3, 4$ . Evaluating residues at these poles and using residue theorem, the density in this case is obtained (for  $0 < v < 1$ ) as

$$\begin{aligned}
 f(v) &= \frac{K(n, 3)}{\sqrt{\pi}} v^{n-7} \left[ \frac{2}{315} v - \frac{4}{45} v^2 - \frac{1}{\sqrt{\pi}} \sum_{r=5}^{\infty} \frac{\Gamma(r-\frac{7}{2})}{r!(r-2)(r-3)(r-4)} v^{r+1} \right. \\
 &\quad \left. + \frac{1}{3} \left( \frac{8}{3} - \ln \frac{v}{4} \right) v^3 - \frac{1}{3} \left( \frac{1}{6} + \ln \frac{v}{4} \right) v^4 + \frac{1}{48} \left( \frac{43}{12} - \ln \frac{v}{4} \right) v^5 \right].
 \end{aligned}$$

For  $p = 4$ , (2.4) reduces to

$$f(v) = K(n, 4)v^{n-9}(2\pi\iota)^{-1} \int_{L_1} \frac{\Gamma(t+1)\Gamma(t+3)}{\prod_{j=5}^7 (t+j)\Gamma(t+\frac{15}{2})\Gamma(t+\frac{13}{2})} v^{-t} dt. \quad (2.7)$$

The integrand has simple poles at  $t = -1, -2$ , poles of order two at  $t = -1 - r$ ,  $r = 2, 3, 7, 8, \dots$ , and poles of order three at  $t = -1 - r$ ,  $r = 4, 5, 6$ . Evaluating residues at these poles and using the residue theorem, we get the density for  $p = 4$  as

$$\begin{aligned}
 f(v) = & K(n, 4)v^{n-9} \left[ \frac{1}{660 \Gamma^2(\frac{11}{2})} v - \frac{1}{270 \Gamma^2(\frac{9}{2})} v^2 \right. \\
 & - \frac{1}{168 \Gamma^2(\frac{7}{2})} \left( \frac{2521}{420} + \ln \frac{v}{16} \right) v^3 - \frac{1}{90 \Gamma^2(\frac{5}{2})} \left( \frac{71}{15} + \ln \frac{v}{16} \right) v^4 \\
 & + \frac{1}{\pi^2} \sum_{r=7}^{\infty} \left\{ \psi(r+1) + \psi(r-1) + \sum_{j=4}^7 \frac{1}{r-j} - \psi(r - \frac{11}{2}) - \psi(r - \frac{9}{2}) - \ln v \right\} \\
 & \frac{\Gamma(r - \frac{9}{2})\Gamma(r - \frac{11}{2})}{r!(r-2)!(r-4)(r-5)(r-6)} v^{r+1} + \frac{1}{12\pi} \left\{ \left( \ln \frac{v}{16} + \frac{13}{4} \right)^2 + \frac{1781}{144} - \frac{2\pi^2}{3} \right\} \frac{v^5}{2} \\
 & - \frac{1}{360\pi} \left\{ \left( \frac{127}{60} - \ln \frac{v}{16} \right)^2 + \frac{31769}{3600} - \frac{2\pi^2}{3} \right\} \frac{v^6}{2} \\
 & \left. - \frac{15}{2(6!)^2\pi} \left\{ \left( -\ln \frac{v}{16} + \frac{121}{30} \right)^2 + \frac{33}{200} - \frac{2\pi^2}{3} \right\} \frac{v^7}{2} \right], 0 < v < 1.
 \end{aligned}$$

For  $p = 5$ , (2.4) slides to

$$\begin{aligned}
 f(v) = & K(n, 5)v^{n-11}(2\pi\iota)^{-1} \int_{L_1} \frac{\Gamma(t+1)\Gamma(t+3)}{\prod_{j=5}^9 (t+j)^{a_j} \Gamma(t + \frac{17}{2})\Gamma(t + \frac{19}{2})} v^{-t} dt, \\
 & 0 < v < 1, \quad (2.8)
 \end{aligned}$$

where  $a_5 = a_6 = a_8 = a_9 = 1$  and  $a_7 = 2$ . The integrand has simple poles at  $t = -1, -2$ , poles of order 2 at  $t = -1 - r$ ,  $r = 2, 3, 9, 10, \dots$ , poles of order 3 at  $t = -5, -6, -8, -9$  and a pole of order 4 at  $t = -7$ . Evaluating residues at

these poles and using residue theorem, the density is derived as

$$\begin{aligned}
 f(v) = & K(n, 5)v^{n-11} \left[ A_1^{(0)}v + A_2^{(0)}v^2 + (-\ln v + B_3^{(0)})A_3^{(0)}v^3 \right. \\
 & + (-\ln v + B_4^{(0)})A_4^{(0)}v^4 + \sum_{j=5,6,8} \left\{ (-\ln v + B_j^{(0)})^2 + B_j^{(1)} \right\} A_j^{(0)} \frac{v^j}{2} \\
 & + \left\{ (-\ln v + B_7^{(0)})^3 + 3(-\ln v)B_7^{(1)} + 3B_7^{(0)}B_7^{(1)} + B_7^{(2)} \right\} A_7^{(0)} \frac{v^7}{3!} \\
 & \left. + \sum_{r=9}^{\infty} (-\ln v + B_r^{(0)})A_r^{(0)}v^r \right], \quad 0 < v < 1.
 \end{aligned}$$

where

$$\begin{aligned}
 A_1^{(0)} &= \frac{12}{(10)!\Gamma^2(\frac{15}{2})}, & A_2^{(0)} &= \frac{4}{5(7)!\Gamma^2(\frac{13}{2})}, & A_3^{(0)} &= \frac{1}{44(6)!\Gamma^2(\frac{11}{2})}, \\
 B_3^{(0)} &= -\frac{5219}{693} + 2 \ln 4, & A_4^{(0)} &= \frac{2}{3(6)!\Gamma^2(\frac{9}{2})}, & B_4^{(0)} &= -\frac{1691}{252} + 2 \ln 4 \\
 A_5^{(0)} &= \frac{5}{(8)!\Gamma^2(\frac{7}{2})}, & B_5^{(0)} &= -\frac{569}{105} + 2 \ln 4, & B_5^{(1)} &= \frac{1202849}{88200} - \frac{2\pi^2}{3}, \\
 A_6^{(0)} &= \frac{2}{15(6)!\Gamma^2(\frac{5}{2})}, & B_6^{(0)} &= -\frac{69}{20} + 2 \ln 4, & B_6^{(1)} &= \frac{10969}{720} - \frac{2\pi^2}{3}, \\
 A_8^{(0)} &= -\frac{1}{3(5)!(7)!\pi}, & B_8^{(0)} &= \frac{989}{210} + 2 \ln 4, & B_8^{(1)} &= \frac{911681}{88200} - \frac{2\pi^2}{3}, \\
 A_7^{(0)} &= \frac{7}{2(9)!\Gamma^2(\frac{3}{2})}, & B_7^{(0)} &= -\frac{2}{15} + 2 \ln 4, & B_7^{(1)} &= \frac{24947}{1800} - \frac{2\pi^2}{3}, \\
 B_7^{(2)} &= -\frac{656261}{54000} + 2\pi^2 - 4\zeta(3), \\
 A_r^{(0)} &= -\frac{\Gamma(r - \frac{13}{2})\Gamma(r - \frac{15}{2})}{\pi^2 r!(r-2)! \prod_{j=5}^9 (r+1-j)^{a_j}}
 \end{aligned}$$

and

$$B_r^{(0)} = \psi(r+1) + \psi(r-1) + \sum_{j=5}^9 a_j (r+1-j)^{-1} - \psi(r - \frac{13}{2}) - \psi(r - \frac{15}{2}).$$



The method employed above can also be used to derive the exact density for  $p \geq 6$ . However, because of the higher order of poles, the expressions for the density will involve generalized Riemann zeta function (see [9]).

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## References

- [1] Amey, A. K. A., Gupta, A. K. *Testing sphericity under a mixture model*, Australian Journal of Statistics, **34** (1992), 451–460.
- [2] Chao, C. C., Gupta, A. K. *Testing of homogeneity of diagonal blocks with blockwise independence*, Communications in Statistics–Theory and Methods, **20**(7) (1991), 1957–1969.
- [3] Gupta, A. K. *On the distribution of sphericity test criterion in the multivariate Gaussian distribution*, Australian Journal of Statistics, **19** (1977), 202–205.
- [4] Gupta, A. K., Chao, C. C. *Nonnull distribution of the likelihood ratio criterion for testing the equality of diagonal blocks with blockwise independence*, Acta Mathematica Scientia, **14**(2) (1994), 195–203.
- [5] Gupta, A. K., Nagar, D. K. *Matrix Variate Distributions*, Chapman & Hall/CRC Press, Boca Raton, 1999.
- [6] Khatri, C. G., Srivastava, M. S. *On exact nonnull distribution of likelihood ratio criteria for sphericity test and equality of two covariance matrices*, Sankhyā, **A33** (1971), 201–206.
- [7] Khatri, C. G., Srivastava, M. S. *Asymptotic expansion of the nonnull distribution of likelihood ratio criteria for covariance matrices*, Annals of Statistics, **2** (1974), 109–117.
- [8] Luke, Y. L. *The Special Functions and Their Approximations*, Vol. I, Academic Press, New York, 1969.
- [9] Magnus, W., Oberhettinger, F., Soni, R. P. *Formulas and Theorems for the Special Functions of Mathematical Physics*, **52**, Springer-Verlag, 1966.

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- [10] Muirhead, R. J. *Forms for distribution for ellipticity statistic for bivariate sphericity*, Communications in Statistics, **A5** (1976), 413–420.
- [11] Nagar, D. K., Jain, S. K., Gupta, A. K. *Distribution of LRC for testing sphericity structure of a covariance matrix in multivariate normal distribution*, Metron, **49** (1991), 435–457.
- [12] Nagar, D. K., Jain, S. K., Gupta, A. K. *On testing circular stationarity and related models*, Journal of Statistical Computation and Simulation, **29** (1988), 225–239.
- [13] Olkin, I. *Testing and estimation for structures which are circular symmetric in blocks*, in *Multivariate Statistical Inference* (D. G. Kabe and R. P. Gupta, eds.), North-Holland, 1973, 183–195.
- [14] Springer, M. D. *Algebra of Random Variables*, John Wiley & Sons, New York, 1979.
- [15] Sugira, N. *Asymptotic expansion of the distribution of likelihood ratio criteria for covariance matrices*, Annals of Mathematical Statistics, **42** (1969), 764–767.
- [16] Sugira, N. *Exact nonnull distributions of sphericity tests for trivariate normal population with power comparison*, American Journal of Mathematical and Management Sciences, **15** (3&4) (1995), 355–374.
- [17] Thomas, D. Roland. *Univariate repeated measures techniques applied to multivariate data*, Psychometrika, **48**(3) (1983), 451–464.