Hindawi Publishing Corporation Journal of Probability and Statistics Volume 2011, Article ID 741701, 11 pages doi:10.1155/2011/741701

Research Article

Large Deviations in Testing Squared Radial Ornstein-Uhleneck Model

Shoujiang Zhao and Qiaojing Liu

College of Science, China Three Gorges University, Yichang 443002, China

Correspondence should be addressed to Shoujiang Zhao, shjzhao@yahoo.com.cn

Received 12 May 2011; Accepted 10 July 2011

Academic Editor: Mohammad Fraiwan Al-Saleh

Copyright © 2011 S. Zhao and Q. Liu. This is an open access article distributed under the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

We study the large deviations and moderate deviations of hypothesis testing for squared radial Ornstein-Uhleneck model. Large deviation principles for the log-likelihood ratio are obtained, by which we give negative regions in testing squared radial Ornstein-Uhleneck model and get the decay rates of the error probabilities.

1. Introduction

Let us consider the hypothesis testing for the following squared radial Ornstein-Uhleneck model:

$$dX_t = (\delta + 2\alpha X_t)dt + 2\sqrt{X_t}dW_t, \quad X_0 = 0, \tag{1.1}$$

where $\alpha < 0$ is the unknown parameter to be tested on the basis of continuous observation of the process $\{X_t, t \geq 0\}$ on the time interval [0,T], W is a standard Brownian motion and, $\delta > 0$ is known. We denote the distribution of the solution (1.1) by P_{α}^{δ} .

We decide the two hypothesis:

$$H_0: \alpha = \alpha_0, \qquad H_1: \alpha = \alpha_1, \tag{1.2}$$

where α_0 , α_1 < 0. The hypothesis testing is based on a partition of

$$\Omega_T = \left\{ \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in \mathbb{R} \right\}$$
(1.3)

of the outcome process on [0,T] into two (decision) regions B_T and its compliment B_T^c , and we decide that H_0 is true or false according to the outcome $X \in B_T$ or $X \in B_T^c$.

The probability $e_1(T)$ of accepting H_1 when H_0 is actually true is called the error probability of the first kind. The probability $e_2(T)$ of accepting H_0 when H_1 is actually true is called the error probability of the second kind. That is,

$$e_1(T) = P_{\alpha_0}^{\delta}(B_T), \qquad e_2(T) = P_{\alpha_1}^{\delta}(B_T^c).$$
 (1.4)

By the Neyman-Pearson lemma (cf. [1]), the optional decision region B_T has the following form:

$$\left\{ \frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \ge c \right\},$$
(1.5)

where $\mathcal{F}_{[0,T]}$ is the σ -algebra generated by the outcome process on [0,T].

The research of hypothesis testing problem has started in the 1930s (cf. [1]). Since the optional decision region B_T has the above form, we are interested in the calculation or approximation of the constant c, and the hypothesis testing problem can be studied by large deviations (cf. [2–6]). In those papers, some large deviation estimates of the error probabilities for some i.i.d. sequences, Markov chains, stationary Gaussian processes, stationary diffusion processes, Ornstein-Uhlenbeck processes are obtained. In this paper, we study the large deviations and moderate deviations for the hypothesis testing problem of squared radial Ornstein-Uhleneck model; by large deviation principle, we obtain that the decay of the error probability of the second kind approaches to 0 or 1 exponentially fast depending on the fixed exponent of the decay of the error probability of the first kind; we also give negative regions and get the decay rates of the error probabilities by moderate deviation principle. The large and moderate deviations for parameter estimators of squared radial Ornstein-Uhleneck model were studied in [7, 8].

2. Main Results

In this section, we state our main results.

Theorem 2.1. Let a(T) be a positive function satisfying

$$\frac{a(T)}{T} \longrightarrow 0, \quad \frac{a(T)}{\sqrt{T}} \longrightarrow \infty \quad as \ T \longrightarrow \infty.$$
 (2.1)

For any a > 0, set

$$B_{T} = \left\{ \frac{1}{a(T)} \log \left. \frac{dP_{\alpha_{1}}^{\delta}}{dP_{\alpha_{0}}^{\delta}} \right|_{\varphi_{[0,T]}} \ge \frac{T}{4a(T)} \frac{(\alpha_{1} - \alpha_{0})^{2} \delta}{\alpha_{0}} - \frac{|\alpha_{1}^{2} - \alpha_{0}^{2}|}{2\alpha_{0}} \sqrt{\frac{a\delta}{-\alpha_{0}}} \right\}. \tag{2.2}$$

Then

$$\lim_{T \to \infty} \frac{T}{a^{2}(T)} \log P_{\alpha_{0}}^{\delta}(B_{T}) = -a,$$

$$\lim_{T \to \infty} \frac{T}{a^{2}(T)} \log P_{\alpha_{1}}^{\delta}(B_{T}^{c}) = -\infty.$$
(2.3)

Theorem 2.2. If $\alpha_1 < \alpha_0 < 0$, then for each a > 0, there exists a $\xi(a) \in \mathbb{R}$, such that

$$\lim_{T \to \infty} \frac{1}{T} \log P_{\alpha_0}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \ge \xi(a) \right) = -a, \tag{2.4}$$

and when $a \in (0, z_{\alpha_1})$,

$$\liminf_{T \to \infty} \frac{1}{T} \log P_{\alpha_1}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} < \xi(a) \right) \le -I_{\alpha_1}(\xi(a)), \tag{2.5}$$

when $a \in (z_{\alpha_1}, +\infty)$,

$$\liminf_{T \to \infty} \frac{1}{T} \log \left(1 - P_{\alpha_1}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} < \xi(a) \right) \right) \le -I_{\alpha_1}(\xi(a)), \tag{2.6}$$

where

$$z_{\alpha_{1}} = -\frac{(\alpha_{0} - \alpha_{1})^{2} \delta}{4\alpha_{1}},$$

$$I_{\alpha_{1}}(z) = \begin{cases} \frac{\alpha_{0}^{2} - \alpha_{1}^{2}}{z + (\alpha_{1} - \alpha_{0})\delta/2} \left(\frac{\delta}{4} - \frac{\alpha_{1}(z + (\alpha_{1} - \alpha_{0})\delta/2)}{\alpha_{1}^{2} - \alpha_{0}^{2}}\right)^{2}, & z < \frac{(\alpha_{0} - \alpha_{1})\delta}{2}, \\ +\infty, & otherwise. \end{cases}$$
(2.7)

Theorem 2.3. If $\alpha_0 < \alpha_1 < 0$, then for each a > 0, there exists a $\tilde{\xi}$ $(a) \in \mathbb{R}$, such that

$$\lim_{T \to \infty} \frac{1}{T} \log P_{\alpha_0}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \ge \widetilde{\xi}(a) \right) = -a, \tag{2.8}$$

and when $a \in (0, \hat{z}_{\alpha_1})$,

$$\liminf_{T \to \infty} \frac{1}{T} \log P_{\alpha_1}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} < \widetilde{\xi}(a) \right) \le -\widehat{I}_{\alpha_1} \left(\widetilde{\xi}(a) \right), \tag{2.9}$$

when $a \in (\hat{z}_{\alpha_1}, +\infty)$,

$$\lim_{T \to \infty} \inf_{\overline{T}} \frac{1}{T} \log \left(1 - P_{\alpha_1}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} < \widetilde{\xi}(a) \right) \right) \le -\widehat{I}_{\alpha_1} \left(\widetilde{\xi}(a) \right), \tag{2.10}$$

where

$$\widehat{z}_{\alpha_{1}} = -\frac{(\alpha_{0} - \alpha_{1})^{2} \delta}{4\alpha_{1}},$$

$$\widehat{I}_{\alpha_{1}}(z) = \begin{cases} \frac{\alpha_{0}^{2} - \alpha_{1}^{2}}{z + (\alpha_{1} - \alpha_{0})\delta/2} \left(\frac{\delta}{4} - \frac{\alpha_{1}(z + (\alpha_{1} - \alpha_{0})\delta/2)}{\alpha_{1}^{2} - \alpha_{0}^{2}}\right)^{2}, & z > \frac{(\alpha_{0} - \alpha_{1})\delta}{2}, \\ +\infty, & otherwise. \end{cases}$$
(2.11)

3. Moderate Deviations in Testing Squared Radial Ornstein-Uhleneck Model

In this section, we will prove Theorem 2.1. Let us introduce the log-likelihood ratio process of squared radial Ornstein-Uhleneck model and study the moderate deviations of the log-likelihood ratio process.

By [7], the log-likelihood ratio process has the representation

$$\log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} = \frac{1}{2} (\alpha_1 - \alpha_0)(X_t - \delta t) - \frac{\alpha_1^2 - \alpha_0^2}{2} \int_0^t X_s^2 ds. \tag{3.1}$$

The following Lemma (cf. [9]) plays an important role in this paper.

Lemma 3.1. The law of X_t under P_{α}^{δ} is $\gamma(\delta/2, \alpha/e^{2t\alpha-1})$, where $\gamma(a, b)$ denotes the Gamma distribution:

$$\gamma(a,b)(dx) = \frac{b^a x^{a-1}}{\Gamma(a)} e^{-bx}(dx), \quad x > 0.$$
 (3.2)

Moreover, for any $\theta \in \mathbb{R}$ *,*

$$\mathbb{E}_{\alpha}^{\delta}\left(e^{\theta X_{t}}\right) = \left(1 - \frac{\theta}{\alpha}\left(e^{2t\alpha} - 1\right)\right)^{-\delta/2}.$$
(3.3)

Lemma 3.2. *For any closed subset* $F \subset \mathbb{R}$ *,*

$$\limsup_{T \to \infty} \frac{T}{a^{2}(T)} \log P_{\alpha_{0}}^{\delta} \left(\frac{1}{a(T)} \left(\log \left. \frac{dP_{\alpha_{1}}^{\delta}}{dP_{\alpha_{0}}^{\delta}} \right|_{\mathcal{F}_{[0,T]}} - \frac{(\alpha_{1} - \alpha_{0})^{2} \delta T}{4\alpha_{0}} \right) \in F \right) \leq -\inf_{z \in F} \frac{-4\alpha_{0}^{3} z^{2}}{\left(\alpha_{1}^{2} - \alpha_{0}^{2}\right)^{2} \delta'}, \tag{3.4}$$

and for any open subset $G \subset \mathbb{R}$,

$$\liminf_{T \to \infty} \frac{T}{a^2(T)} \log P_{\alpha_0}^{\delta} \left(\frac{1}{a(T)} \left(\log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} - \frac{(\alpha_1 - \alpha_0)^2 \delta T}{4\alpha_0} \right) \in G \right) \ge -\inf_{z \in G} \frac{-4\alpha_0^3 z^2}{\left(\alpha_1^2 - \alpha_0^2\right)^2 \delta}.$$
(3.5)

Proof. Let

$$\Lambda_{T}(y) = \log \mathbb{E}_{\alpha_{0}}^{\delta} \exp \left\{ \frac{a(T)y}{T} \left(\log \left. \frac{dP_{\alpha_{1}}^{\delta}}{dP_{\alpha_{0}}^{\delta}} \right|_{\varphi_{[0,T]}} - \frac{(\alpha_{1} - \alpha_{0})^{2} \delta T}{4\alpha_{0}} \right) \right\}. \tag{3.6}$$

By (3.1), for any φ < 0, we have

$$\Lambda_{T}(y) = -\frac{a(T)y\delta(\alpha_{1} - \alpha_{0})^{2}}{4\alpha_{0}} + \log \mathbb{E}_{\alpha_{0}}^{\delta} \exp\left\{\lambda(X_{T} - \delta T) + u \int_{0}^{T} X_{t}^{2} ds\right\}$$

$$= -\frac{a(T)y\delta(\alpha_{1} - \alpha_{0})^{2}}{4\alpha_{0}} + \log \mathbb{E}_{\varphi}^{\delta} \left[\frac{dP_{\alpha_{0}}^{\delta}}{dP_{\varphi}^{\delta}} \exp\left\{\lambda(X_{T} - \delta T) + u \int_{0}^{T} X_{t}^{2} ds\right\}\right]$$

$$= -\frac{a(T)y\delta(\alpha_{1}^{2} - \alpha_{0}^{2})}{4\alpha_{0}} + \log \mathbb{E}_{\varphi}^{\delta} \left[\exp\left\{\left(\lambda + \frac{\alpha_{0} - \varphi}{2}\right)X_{T} - \frac{\alpha_{0} - \varphi}{2}\delta T + \left(u - \frac{1}{2}\alpha_{0}^{2} + \frac{1}{2}\varphi^{2}\right)\int_{0}^{T} X_{s}^{2} ds\right\}\right], \tag{3.7}$$

where

$$\lambda = \frac{a(T)y(\alpha_1 - \alpha_0)}{2T}, \qquad u = -\frac{a(T)y(\alpha_1^2 - \alpha_0^2)}{2T}.$$
 (3.8)

For *T* large enough, $\alpha_0^2 - 2u > 0$, we can choose $\varphi = -\sqrt{\alpha_0^2 - 2u}$, then $\varphi < 0$ and

$$\Lambda_T(y) = -\frac{a(T)y\delta(\alpha_1^2 - \alpha_0^2)}{4\alpha_0} - \frac{\alpha_0 - \varphi}{2}\delta T + \log \mathbb{E}_{\varphi}^{\delta} \left[\exp\left\{ \left(\lambda + \frac{\alpha_0 - \varphi}{2} \right) X_T \right\} \right]. \tag{3.9}$$

By Lemma 3.1, we have

$$\lim_{T \to \infty} \frac{T}{a^{2}(T)} \log \mathbb{E}_{\varphi}^{\delta} \left[\exp \left\{ \left(\lambda + \frac{\alpha_{0} - \varphi}{2} \right) X_{T} \right\} \right]$$

$$= \lim_{T \to \infty} -\frac{T\delta}{2a^{2}(T)} \log \left(1 - \frac{1}{\varphi} \left(\lambda + \frac{\alpha_{0} - \varphi}{2} \right) \left(e^{2T\varphi} - 1 \right) \right) = 0.$$
(3.10)

Therefore,

$$\Lambda(y) := \lim_{T \to \infty} \frac{T}{a^{2}(T)} \Lambda_{T}(y)
= \lim_{T \to \infty} \frac{T}{a^{2}(T)} \left(-\frac{a(T)y\delta(\alpha_{1}^{2} - \alpha_{0}^{2})}{4\alpha_{0}} - \frac{\alpha_{0} - \varphi}{2} \delta T \right)
= \lim_{T \to \infty} \frac{T}{a^{2}(T)} \left(-\frac{a(T)y\delta(\alpha_{1}^{2} - \alpha_{0}^{2})}{4\alpha_{0}} - \frac{\alpha_{0}\delta T}{2} \left(1 - \sqrt{1 + \frac{a(T)y(\alpha_{1}^{2} - \alpha_{0}^{2})}{T\alpha_{0}^{2}}} \right) \right)
= \frac{y^{2}}{16} \frac{(\alpha_{1}^{2} - \alpha_{0}^{2})^{2} \delta}{-\alpha_{0}^{3}}.$$
(3.11)

Finally, the Gärtner-Ellis theorem (cf. [10]) implies the conclusion of Lemma 3.2.

Noting that

$$\log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} = -\log \left. \frac{dP_{\alpha_0}^{\delta}}{dP_{\alpha_1}^{\delta}} \right|_{\varphi_{[0,T]}},\tag{3.12}$$

we also have the following result.

Lemma 3.3. *For any closed subset* $F \subset \mathbb{R}$ *,*

$$\limsup_{T \to \infty} \frac{T}{a^{2}(T)} \log P_{\alpha_{1}}^{\delta} \left(\frac{1}{a(T)} \left(\log \left. \frac{dP_{\alpha_{1}}^{\delta}}{dP_{\alpha_{0}}^{\delta}} \right|_{\mathcal{F}_{[0,T]}} + \frac{(\alpha_{1} - \alpha_{0})^{2} \delta T}{4\alpha_{1}} \right) \in F \right) \leq -\inf_{z \in F} \frac{-4\alpha_{1}^{3} z^{2}}{\left(\alpha_{1}^{2} - \alpha_{0}^{2}\right)^{2} \delta'}$$
(3.13)

and for any open subset $G \subset \mathbb{R}$,

$$\lim_{T \to \infty} \inf \frac{T}{a^{2}(T)} \log P_{\alpha_{1}}^{\delta} \left(\frac{1}{a(T)} \left(\log \left. \frac{dP_{\alpha_{1}}^{\delta}}{dP_{\alpha_{0}}^{\delta}} \right|_{\varphi_{[0,T]}} + \frac{(\alpha_{1} - \alpha_{0})^{2} \delta T}{4\alpha_{1}} \right) \in G \right) \ge -\inf_{z \in G} \frac{-4\alpha_{1}^{3} z^{2}}{\left(\alpha_{1}^{2} - \alpha_{0}^{2}\right)^{2} \delta}.$$
(3.14)

Proof of Theorem 2.1. The first claim is a direct conclusion of Lemma 3.2. Since

$$\frac{T}{a(T)} \frac{(\alpha_1 - \alpha_0)^2 \delta}{\alpha_0} + \frac{T}{a(T)} \frac{(\alpha_1 - \alpha_0)^2 \delta}{\alpha_1} \longrightarrow -\infty, \quad \text{as } T \longrightarrow \infty, \tag{3.15}$$

by Lemma 3.3, we see that the second one also holds.

4. Large Deviations in Testing Fractional Ornstein-Uhleneck Model

In this section, we will prove Theorems 2.2 and 2.3. We first study the large deviations of the log-likelihood ratio process.

Lemma 4.1. Assume $\alpha_1 < \alpha_0 < 0$. Then for any closed subset $F \subset \mathbb{R}$,

$$\limsup_{T \to \infty} \frac{1}{T} \log P_{\alpha_0}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in F \right) \le -\inf_{z \in F} I_{\alpha_0}(z), \tag{4.1}$$

and for any open subset $G \subset \mathbb{R}$,

$$\liminf_{T \to \infty} \frac{1}{T} \log P_{\alpha_0}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in G \right) \ge -\inf_{z \in G} I_{\alpha_0}(z), \tag{4.2}$$

where

$$I_{\alpha_{0}}(z) = \begin{cases} \frac{\alpha_{0}^{2} - \alpha_{1}^{2}}{z + (\alpha_{1} - \alpha_{0})\delta/2} \left(\frac{\delta}{4} - \frac{\alpha_{0}(z + (\alpha_{1} - \alpha_{0})\delta/2)}{\alpha_{1}^{2} - \alpha_{0}^{2}} \right)^{2}, & z < \frac{(\alpha_{0} - \alpha_{1})\delta}{2}, \\ +\infty, & otherwise. \end{cases}$$
(4.3)

Proof. Let

$$\Lambda_T(y) = \log \mathbb{E}_{\alpha_0}^{\delta} \exp \left\{ y \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \right\}. \tag{4.4}$$

Then for φ < 0, we have

$$\Lambda_{T}(y) = \log \mathbb{E}_{\alpha_{0}}^{\delta} \exp \left\{ \lambda(X_{T} - \delta T) + u \int_{0}^{T} X_{t}^{2} ds \right\}$$

$$= \log \mathbb{E}_{\varphi}^{\delta} \left[\frac{dP_{\alpha_{0}}^{\delta}}{dP_{\varphi}^{\delta}} \exp \left\{ \lambda(X_{T} - \delta T) + u \int_{0}^{T} X_{t}^{2} ds \right\} \right]$$

$$= \log \mathbb{E}_{\varphi}^{\delta} \left[\exp \left\{ \left(\lambda + \frac{\alpha_{0} - \varphi}{2} \right) (X_{T} - \delta T) + \left(u - \frac{1}{2}\alpha_{0}^{2} + \frac{1}{2}\varphi^{2} \right) \int_{0}^{T} X_{s}^{2} ds \right\} \right], \tag{4.5}$$

where $\lambda = (\alpha_1 - \alpha_0)y/2$, $u = y(\alpha_0^2 - \alpha_1^2)/2$.

Since $\alpha_0^2 - 2u > 0$, for $y > (\alpha_0^2/\alpha_0^2 - \alpha_1^2)$, we can choose $\varphi = -\sqrt{\alpha_0^2 - 2u}$, for each $y > (\alpha_0^2/\alpha_0^2 - \alpha_1^2)$; then $\varphi < 0$ and

$$\Lambda_T(y) = -\delta T \left(\lambda + \frac{\alpha_0 - \varphi}{2} \right) + \log \mathbb{E}_{\varphi}^{\delta} \left[e^{(\lambda + (\alpha_0 - \varphi)/2)X_T} \right]. \tag{4.6}$$

By Lemma 3.1, we get

$$\lim_{T \to \infty} \frac{1}{T} \log \mathbb{E}_{\varphi}^{\delta} \left[\exp \left\{ \left(\lambda + \frac{\alpha_0 - \varphi}{2} \right) X_T \right\} \right] = 0. \tag{4.7}$$

Therefore,

$$\Lambda(y) =: \lim_{T \to \infty} \frac{1}{T} \Lambda_T(y) = -\frac{\delta}{2} \left((\alpha_1 - \alpha_0)y + \alpha_0 + \sqrt{\alpha_0^2 + (\alpha_1^2 - \alpha_0^2)y} \right). \tag{4.8}$$

Since $\Lambda(y)$ is a strictly convex differentiable function on $\mathfrak{D}_{\Lambda} = (\alpha_{-}, +\infty)$ with

$$\alpha_{-} = \frac{\alpha_0^2}{\alpha_0^2 - \alpha_1^2} < 0,$$

$$\lim_{y \to \alpha_{-}} \Lambda'(y) = +\infty,$$
(4.9)

where \mathfrak{D}_{Λ} is the effective domain of Λ , we see that $\Lambda(y)$ is steep. Finally, by

$$\sup_{y \in \mathbb{R}} \left\{ zy - \Lambda(y) \right\} = \begin{cases} \frac{\alpha_0^2 - \alpha_1^2}{z + (\alpha_1 - \alpha_0)\delta/2} \left(\frac{\delta}{4} - \frac{\alpha_0(z + (\alpha_1 - \alpha_0)\delta/2)}{\alpha_1^2 - \alpha_0^2} \right)^2, & z < \frac{(\alpha_0 - \alpha_1)\delta}{2}, \\ +\infty, & \text{otherwise,} \end{cases}$$

$$(4.10)$$

and Gärtner-Ellis theorem, we complete the proof of this lemma.

Similarly, when $\alpha_0 < \alpha_1 < 0$, we have

$$\Lambda(y) =: \lim_{T \to \infty} \frac{1}{T} \Lambda_T(y) = -\frac{\delta}{2} \left((\alpha_1 - \alpha_0)y + \alpha_0 + \sqrt{\alpha_0^2 + (\alpha_1^2 - \alpha_0^2)y} \right). \tag{4.11}$$

Since $\Lambda(y)$ is a strictly convex differentiable function on $\mathfrak{D}_{\Lambda} = (-\infty, \alpha_+)$ with

$$\alpha_{+} = \frac{\alpha_{0}^{2}}{\alpha_{0}^{2} - \alpha_{1}^{2}} > 0, \tag{4.12}$$

and $\lim_{y\to a_+} \Lambda'(y) = +\infty$, we can see that $\Lambda(y)$ is steep. By Gärtner-Ellis theorem, we also have the following result.

Lemma 4.2. Assume $\alpha_0 < \alpha_1 < 0$. Then for any closed subset $F \subset \mathbb{R}$,

$$\limsup_{T \to \infty} \frac{1}{T} \log P_{\alpha_0}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in F \right) \le -\inf_{z \in F} \widehat{I}_{\alpha_0}(z), \tag{4.13}$$

and for any open subset $G \subset \mathbb{R}$,

$$\liminf_{T \to \infty} \frac{1}{T} \log P_{\alpha_0}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in G \right) \ge -\inf_{z \in G} \widehat{I}_{\alpha_0}(z), \tag{4.14}$$

where

$$\widehat{I}_{\alpha_0}(z) = \begin{cases} \frac{\alpha_0^2 - \alpha_1^2}{z + ((\alpha_1 - \alpha_0)\delta/2)} \left(\frac{\delta}{4} - \frac{\alpha_0(z + (\alpha_1 - \alpha_0)\delta/2)}{\alpha_1^2 - \alpha_0^2}\right)^2, & z > \frac{(\alpha_0 - \alpha_1)\delta}{2}, \\ +\infty, & otherwise. \end{cases}$$
(4.15)

Note that

$$\log \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}}\bigg|_{\varphi_{[0,T]}} = -\log \frac{dP_{\alpha_0}^{\delta}}{dP_{\alpha_1}^{\delta}}\bigg|_{\varphi_{[0,T]}}.$$
(4.16)

Then we have the following Lemma.

Lemma 4.3. Assume $\alpha_1 < \alpha_0 < 0$. Then for any closed subset $F \subset \mathbb{R}$,

$$\limsup_{T \to \infty} \frac{1}{T} \log P_{\alpha_1}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in F \right) \le -\inf_{z \in F} \widehat{I}_{\alpha_1}(z), \tag{4.17}$$

and for any open subset $G \subset \mathbb{R}$,

$$\liminf_{T \to \infty} \frac{1}{T} \log P_{\alpha_1}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in G \right) \ge -\inf_{z \in G} \widehat{I}_{\alpha_1}(z), \tag{4.18}$$

where

$$I_{\alpha_{1}}(z) = \begin{cases} \frac{\alpha_{0}^{2} - \alpha_{1}^{2}}{z + (\alpha_{1} - \alpha_{0})\delta/2} \left(\frac{\delta}{4} - \frac{\alpha_{1}(z + (\alpha_{1} - \alpha_{0})\delta/2)}{\alpha_{1}^{2} - \alpha_{0}^{2}}\right)^{2}, & z < \frac{(\alpha_{0} - \alpha_{1})\delta}{2}, \\ +\infty, & otherwise. \end{cases}$$
(4.19)

Lemma 4.4. Assume $\alpha_0 < \alpha_1 < 0$. Then for any closed subset $F \subset \mathbb{R}$,

$$\limsup_{T \to \infty} \frac{1}{T} \log P_{\alpha_1}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in F \right) \le -\inf_{z \in F} I_{\alpha_1}(z), \tag{4.20}$$

and for any open subset $G \subset \mathbb{R}$,

$$\lim_{T \to \infty} \inf \frac{1}{T} \log P_{\alpha_1}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \in G \right) \ge -\inf_{z \in G} I_{\alpha_1}(z), \tag{4.21}$$

where

$$\widehat{I}_{\alpha_{1}}(z) = \begin{cases} \frac{\alpha_{0}^{2} - \alpha_{1}^{2}}{z + (\alpha_{1} - \alpha_{0})\delta/2} \left(\frac{\delta}{4} - \frac{\alpha_{1}(z + (\alpha_{1} - \alpha_{0})\delta/2)}{\alpha_{1}^{2} - \alpha_{0}^{2}} \right)^{2}, & z > \frac{(\alpha_{0} - \alpha_{1})\delta}{2}, \\ +\infty, & otherwise. \end{cases}$$
(4.22)

By the expression of $I_{\alpha_0}(z)$, $I_{\alpha_1}(z)$, $\widehat{I}_{\alpha_0}(z)$, and $\widehat{I}_{\alpha_1}(z)$, the following lemma is.

Lemma 4.5. (i)

$$I_{\alpha_0}(z) = 0 \quad \text{iff } z_{\alpha_0} = \frac{(\alpha_0 - \alpha_1)^2 \delta}{4\alpha_0},$$

$$I_{\alpha_1}(z) = 0 \quad \text{iff } z_{\alpha_1} = -\frac{(\alpha_1 - \alpha_0)^2 \delta}{4\alpha_1},$$
(4.23)

for all $z < (\alpha_0 - \alpha_1)\delta/2$,

$$I_{\alpha_0}(z) = I_{\alpha_1}(z) + z;$$
 (4.24)

(ii)

$$\widehat{I}_{\alpha_0}(z) = 0 \quad \text{iff } \widehat{z}_{\alpha_0} = \frac{(\alpha_0 - \alpha_1)^2 \delta}{4\alpha_0},$$

$$\widehat{I}_{\alpha_1}(z) = 0 \quad \text{iff } \widehat{z}_{\alpha_1} = -\frac{(\alpha_1 - \alpha_0)^2 \delta}{4\alpha_1},$$
(4.25)

for all $z > (\alpha_0 - \alpha_1)\delta/2$

$$\widehat{I}_{\alpha_0}(z) = \widehat{I}_{\alpha_1}(z) + z. \tag{4.26}$$

Proof of Theorems 2.2 *and* 2.3. Since the proofs of the two theorems are similar, we only prove Theorem 2.2. Since $I_{\alpha_0}(z)$ is increasing on $(z_{\alpha_0}, (\alpha_0 - \alpha_1)\delta/2)$ and $I_{\alpha_0}(z_{\alpha_0}) = 0$, Therefore, for a > 0, by Lemma 4.1, we can choose a $\xi(a) \in (z_{\alpha_0}, (\alpha_0 - \alpha_1)\delta/2))$ such that

$$\lim_{T \to \infty} \frac{1}{T} \log P_{\alpha_0}^{\delta} \left(\frac{1}{T} \log \left. \frac{dP_{\alpha_1}^{\delta}}{dP_{\alpha_0}^{\delta}} \right|_{\varphi_{[0,T]}} \ge \xi(a) \right) = -a. \tag{4.27}$$

It is clear that $\xi(a)$ is increasing for a > 0, and by Lemma 4.5, we get $I_{\alpha_0}(z_{\alpha_1}) = z_{\alpha_1}$, which implies $\xi(z_{\alpha_1}) = z_{\alpha_1}$. Hence for $a \in (0, z_{\alpha_1})$, we have $\xi(a) \in (0, z_{\alpha_1})$, and since $I_{\alpha_1}(z)$ is nonincreasing for $z \le \xi(a)$, therefore we get

$$I_{\alpha_1}(\xi(a)) = \inf\{I_{\alpha_1}(z) : z \le \xi(a)\}. \tag{4.28}$$

Similarly, for $a \in (z_{\alpha_1}, +\infty)$, we have $\xi(a) \in (z_{\alpha_1}, (\alpha_0 - \alpha_1)\delta/2)$, and since $I_{\alpha_1}(z)$ is non-decreasing for $z \ge \xi(a)$, therefore we get

$$I_{\alpha_1}(\xi(a)) = \inf\{I_{\alpha_1}(z) : z \ge \xi(a)\},\tag{4.29}$$

which complete the proof of Theorem 2.2.

Acknowledgments

The authors would like to express their gratitude to Professor F. Q. Gao and the reviewer for their valuable comments.

References

- [1] J. Neyman and E. S. Pearson, "On the problem of the most efficient tests of statistical hypotheses," *Philosophical Transactions of the Royal Society of London*, vol. 231, pp. 289–337, 1933.
- [2] R. E. Blahut, "Hypothesis testing and information theory," *IEEE Transactions on Information Theory*, vol. 20, pp. 405–417, 1984.
- [3] T. Chiyonobu, "Hypothesis testing for signal detection problem and large deviations," Nagoya Mathematical Journal, vol. 162, pp. 187–203, 2003.
- [4] P. V. Gapeev and U. Küchler, "On large deviations in testing Ornstein-Uhlenbeck-type models," *Statistical Inference for Stochastic Processes*, vol. 11, no. 2, pp. 143–155, 2008.
- [5] T. S. Han and K. Kobayashi, "The strong converse theorem for hypothesis testing," *IEEE Transactions on Information Theory*, vol. 35, no. 1, pp. 178–180, 1989.
- [6] K. Nakagawa and F. Kanaya, "On the converse theorem in statistical hypothesis testing for Markov chains," *IEEE Transactions on Information Theory*, vol. 39, no. 2, pp. 629–633, 1933.
- [7] M. Zani, "Large deviations for squared radial Ornstein-Uhlenbeck processes," *Stochastic Processes and Their Applications*, vol. 102, no. 1, pp. 25–42, 2002.
- [8] F. Q. Gao and H. Jiang, "Moderate deviations for squared Ornstein-Uhlenbeck process," *Statistics & Probability Letters*, vol. 79, no. 11, pp. 1378–1386, 2009.
- [9] J. Pitman and M. Yor, "A decomposition of Bessel bridges," Zeitschrift für Wahrscheinlichkeitstheorie und Verwandte Gebiete, vol. 59, no. 4, pp. 425–457, 1982.
- [10] A. Dembo and O. Zeitouni, Large Deviation Technique and Applicatons, Springer, New York, NY, USA,

















Submit your manuscripts at http://www.hindawi.com











Journal of Discrete Mathematics











